Tony Ware Joint work with Yu Li

ICCF 2-5 April 2024, Amsterdam

- Scenario: we want to estimate a quantity  $\overline{P} = \mathbb{E}[P]$ .
- We cannot sample P directly. But we do have access to a sequence of *approximate* estimators P<sup>l</sup><sub>l</sub> for l = 0, 1, 2, ....
- ▶  $\overline{P}_l = \mathbb{E}[P_l^l] \to \overline{P}$  as  $l \to \infty$ , but the samples become increasing costly to generate.
- Typically, |P
  <sub>l</sub> P| is inversely proportional to the cost, while the sampling error is inversely proportional to the square of the number of samples, so that the cost of generating an overall error of *ε* is proportional to 1/*ε*<sup>3</sup>.

**MLMC** (Giles, 2008) exploits multiple levels of estimates to reduce the cost to - under certain conditions - as low as  $1/\epsilon^2$ .

- Suppose that, whenever we generate a sample of P<sup>l</sup><sub>l</sub>, we are able to generate a correlated sample of P<sup>l</sup><sub>l-1</sub>, with mean <u>P</u><sub>l-1</sub>.
- We collect these together as basic estimators:

$$Y_0 = P_0^0$$
, and  $Y_l = P_l^l - P_{l-1}^l$  for  $l = 1, 2, ...$ 

• A MLMC estimator at level *L* takes the form  $\Big| \mathcal{P}_L = \sum_{\alpha_l^L} Y_l \Big|$ 



• The numbers of samples  $\alpha_l^L$  are chosen to minimise the cost per unit variance  $W_L^2$ :

$$\alpha_l^L = \frac{\Delta_l}{\eta_l} W_L$$
, with  $W_L = \sum_{l=0}^L \Delta_l \eta_l$ ,

where  $\Delta_l^2$  is the variance of  $Y_l$  and  $\eta_l^2$  is the cost of computing  $Y_l$ .

#### Two ideas

#### First idea

The ratio of successive numbers of samples is *independent of L*, so we can write the MLMC estimator **recursively**:

$$\boxed{\mathcal{P}_0 = {}_{\Delta_0^2}Y_0,} \text{ and, for } l > 0, \quad \mathcal{P}_l = \sum_{l'=0}^l {}_{\alpha_{l'}^l}Y_{l'} = {}_{\alpha_l}Y_l + {}_{\beta_l}\mathcal{P}_{l-1},$$

where 
$$\alpha_l = \alpha_l^l = \frac{\Delta_l W_l}{\eta_l}$$
 and  $\beta_l = \frac{W_l}{W_{l-1}}$ .

Expanding  $\mathcal{P}_{l} = {}_{\alpha_{l}}P_{l}^{l} - ({}_{\alpha_{l}}P_{l-1}^{l} - {}_{\beta_{l}}\mathcal{P}_{l-1})$ , and noticing that  $\mathbb{E}[{}_{\alpha_{l}}P_{l-1}^{l} - {}_{\beta_{l}}\mathcal{P}_{l-1}] = 0$ , leads to ...

Two ideas

#### Second idea

...adding **weights** (in the spirit of control variates (Kebaier 2005)):

$$\mathcal{P}_{l} = {}_{\alpha_{l}}P_{l}^{l} - \theta_{l-1} \left( {}_{\alpha_{l}}P_{l-1}^{l} - {}_{\beta_{l}}\mathcal{P}_{l-1} \right) = {}_{\alpha_{l}}Y_{l}^{\theta_{l-1}} + \theta_{l-1}{}_{\beta_{l}}\mathcal{P}_{l-1}$$

where  $Y_l^{\theta} = P_l^l - \theta P_{l-1}^l$ . This results in the *weighted MLMC estimate*:

$$\mathcal{P}_L = \sum_{l=0}^L \Theta_l^L \left( \frac{\alpha_l \beta_L}{\beta_l} \right) Y_l^{\theta_l},$$

where 
$$\Theta_l^L = \prod_{l'=l}^{L-1} \theta_{l'}$$
.

- We optimise both the numbers of samples and the weights. It is most natural to do this recursively, minimising the cost of P<sub>l</sub> subject to unit variance.
- We start with  $\mathcal{P}_0 = {}_{\alpha_0}Y_0 = {}_{\alpha_0}P_0^0$ . In this case,  $\alpha_0 = \Delta_0^2$ , and  $W_0 = \Delta_0\eta_0$ , so that  $\mathbb{V}[\mathcal{P}_0] = 1$ .
- (Note that we can set the variances to some value v<sup>2</sup> by multiplying all the values of α by v<sup>2</sup>.)
- Suppose we have an optimised WMLMC estimator *P*<sub>*l*-1</sub>.
   For fixed *θ*<sub>*l*-1</sub>, we can minimise the cost by setting

$$\alpha_{l} = \frac{\Delta_{l}^{\theta_{l-1}} W_{l}}{\eta_{l}}, \quad \beta_{l} = |\theta_{l-1}| \frac{W_{l}}{W_{l-1}},$$
$$W_{l} = \Delta_{l}^{\theta_{l-1}} \eta_{l} + |\theta_{l-1}| W_{l-1}.$$

- If  $\rho_l = \operatorname{Corr}[P_l^l, P_{l-1}^l] \leq \frac{W_{l-1}}{\eta_l \sigma_{l-1}}$ , we take  $\theta_{l-1} = 0$ , and l becomes our new coarsest level.
- Otherwise, the optimal  $\theta_{l-1}$  is positive, and we have

$$\theta_{l-1} = \frac{\rho_l \sigma_l}{\sigma_{l-1}} - \frac{\Delta_l^{\theta_{l-1}} W_{l-1}}{\eta_l \sigma_l^2}, \text{ with } \Delta_l^{\theta_{l-1}} = \sqrt{\frac{\sigma_l^2 (1 - \rho_l^2)}{1 - \frac{W_{l-1}^2}{\eta_l^2 \sigma_l^2}}}$$

This results in

$$\frac{W_l}{\sigma_l} = \rho_l \frac{W_{l-1}}{\sigma_{l-1}} + \sqrt{(1 - \rho_l^2) \left(\eta_l^2 - \frac{W_{l-1}^2}{\sigma_{l-1}^2}\right)}$$

which allows us to determine the optimal weights  $\theta_l$  and the efforts  $\alpha_l$  and  $\beta_l$ .

#### Comparison with MLMC

- ▶ In order to gain insight into the comparative the costs associated with MLMC and WMLMC it is convenient to assume that  $\sigma_l \equiv \sigma$  and that  $\eta_l^2 = M \eta_{l-1}^2$ , and to focus on the two-level case. We have that  $W_0 = \sigma \eta_0 = \widehat{W}_0$  (where  $\widehat{W}_l$  is the cost of a single-level estimator at level *l*).
- We have, for the optimal weighted version

$$W_1 = \begin{cases} \widehat{W}_0 \left( \rho_1 + \sqrt{(1 - \rho_1^2)(M - 1)} \right) & \text{if } \rho_1 > \frac{1}{\sqrt{M}} \\ \widehat{W}_1 \ (= \sigma \eta_1) & \text{otherwise.} \end{cases}$$

► And for MLMLC (which corresponds to taking  $\theta_0 = 1$ )

$$W_1 = \begin{cases} \widehat{W}_0 \left( 1 + \sqrt{2M(1-\rho_1)} \right) & \text{if } \rho_1 > 1 - \frac{1}{2} \left( 1 - \frac{1}{\sqrt{M}} \right)^2 \\ \widehat{W}_1 & \text{otherwise.} \end{cases}$$



#### 11

# Weighted Multilevel Monte Carlo



#### 12

# Weighted Multilevel Monte Carlo





#### 14

## Weighted Multilevel Monte Carlo

#### Two-level cost ratio (unweighted/weighted)



#### Three-level cost ratio (unweighted/weighted)











### Weighted Multi-Index MLMC

- Our estimates are now indexed by multi-indices λ ∈ N<sup>2</sup>.
- The MIMLMC method uses alternating signs to define Y<sub>λ</sub>.
- It can (again) be written in a recursive form, and weights can be added.
- There is no longer an explicit expression for the optimal weights, but they can be determined numerically at each node.

$$\begin{array}{c} (0, 2) \\ (1, 2$$

$$\mathcal{P}_{\lambda} = {}_{\alpha_{\lambda}}Y_{\lambda}^{t_{\lambda}} + \sum_{\mu \in \Box_{\lambda}^{-}} \theta_{\mu}^{\lambda} {}_{\left(\frac{\beta_{\lambda}}{\beta_{\mu}}\right)} \mathcal{P}_{\mu}$$
$$= \sum_{\mu \leq \lambda} \Theta_{\mu}^{\lambda} {}_{\left(\frac{\alpha_{\mu}\beta_{\lambda}}{\beta_{\mu}}\right)} Y_{\mu}.$$

Some more details

#### We write

• The index set  $\Box_{\lambda}^{-} = \{\mu \ge (0,0) | \max_i \lambda_i - \mu_i = 1\}.$ 

- $\Theta^{\lambda}_{\mu}$  is the sum over all paths from  $\mu$  to  $\lambda$  of the product of the values of  $\theta^{\lambda'}_{\mu'}$  along each path.
- The effort  $\frac{\alpha_{\mu}\beta_{\lambda}}{\beta_{\mu}}$  is independent of the path from  $\mu$  to  $\lambda$  by virtue of the recursive construction.

Given the recursive relation

$$\mathcal{P}_{\lambda} = {}_{\alpha_{\lambda}}Y_{\lambda}^{t_{\lambda}} + \sum_{\mu \in \Box_{\lambda}^{-}} \theta_{\mu}^{\lambda} \left( {}_{\frac{\beta_{\lambda}}{\beta_{\mu}}} \right) \mathcal{P}_{\mu},$$

we can, given  $t_{\lambda} = [\theta_{\mu}^{\lambda}]_{\mu \in \Box_{\lambda}^{-}}$ , determine  $\alpha_{\lambda}$  and  $\beta_{\lambda}$  so as to minimise the effort needed to ensure  $\mathbb{V}[\mathcal{P}_{\lambda}] = 1$ .

 It remains to optimise over t<sub>λ</sub>, and this involves minimising the sum of the square roots of two quadratic forms defined by t<sub>λ</sub>. This minimisation is performed numerically at each node.

• The unweighted MIMLMC method consists in setting  $\theta^{\lambda}_{\mu} = (-1)^{1+|\lambda-\mu|}$ .

Numerical comparisons: Zakai SPDE

We consider the Zakai SPDE (Reisinger/Wang 2018):

$$\mathrm{d}v = -\mu \frac{\partial v}{\partial x} \mathrm{d}t + \frac{1}{2} \frac{\partial^2 v}{\partial x^2} \mathrm{d}t - \sqrt{\rho} \frac{\partial v}{\partial x} \mathrm{d}M_t.$$

The quantity of interest is  $L_t = 1 - \int_0^\infty v(t, x) dx$ . We compute with a timestep  $k = 4^{-m+1}$  and a grid spacing in x of  $h = 2^{-n}$ , and compare the effort required to achieve a fixed variance using the unweighted and weighted MIMLMC:

$n \backslash m$	0	1	2	3
0	1.0	1.0	1.0	1.0
1	1.6	1.5	1.5	1.7
2	1.8	1.9	1.8	2.0
3	2.0	1.9	1.9	1.9

Numerical comparisons: a hybrid COS-Monte-Carlo basket option valuation

► Here we express a (d+2)-asset basket option payoff as a function  $\Lambda(Z_1, Z_2)$  of independent normal random variables  $Z_1 \in \mathbb{R}^2$  and  $Z_2 \in \mathbb{R}^d$ . The option value can be written as a nested expectation:

$$\mathbb{E}[\Lambda_2(Z_2)], \text{ with } \Lambda_2(z_2) = \mathbb{E}[\Lambda(Z_1, z_2)].$$

- For each sample of  $Z_2$ , the inner expectation is computed using a two-dimensional COS method, with  $(2^{1+L_i} - 1)$ modes used in dimension *i* (for i = 1, 2).
- We again compare MIMC with WMIMC. For some combinations of L<sub>1</sub> and L<sub>2</sub>, MIMC provides no advantage over a single-level estimate.

Tony Ware, April 2024 WMLMC

## Weighted Multi-Index Multilevel Monte Carlo

Numerical comparisons: a hybrid COS-Monte-Carlo basket option valuation



Numerical comparisons: a hybrid COS-Monte-Carlo basket option valuation

The ratio between the unweighted and weighted MIMLMC costs to achieve a given variance

$L_1 \setminus L_2$	0	1	2	3	4	5	6
0							
1	1.49	1.59	1.57	1.56	1.75	1.71	1.70
2	1.46	1.61	1.86	2.85	3.51	3.17	4.26
3	1.59	1.81	1.93	6.18	6.57	11.78	14.58
4	1.58	2.17	4.25	11.93	10.44	15.22	18.47
5	1.61	2.36	5.81	12.39	18.51	24.46	27.37
6	1.64	2.35	5.82	13.25	21.41	28.56	31.48

### Conclusions

- MLMC and MIMLMC can be formulated recursively, and in that way are naturally viewed as nested control variate variance reduction techniques.
- As such, weights can be added, and the optimal weights computed at each node.
- The addition of weights does not change the asymptotic rate of convergence, but it does allow for more efficient use of estimates at coarser resolutions (and lower correlations) than with unweighted MLMC, resulting in potentially significant gains in performance.
- The gains are relatively insensitive to the optimal choice of weights.
- The multi-index version offers potentially even greater relative improvement.

#### References

#### References I

- Amri, M. R. E., Mycek, P., Ricci, S., and De Lozzo, M. (2023). Multilevel surrogate-based control variates. arXiv preprint arXiv:2306.10800.
- Giles, M. B. (2008). Multilevel Monte Carlo path simulation. Operations Research-Baltimore, 56(3):607-617.
- Giles, M. B. (2015). Multilevel Monte Carlo methods. Acta Numerica, 24:259-328.
- Giorgi, D., Lemaire, V., and Pagès, G. (2020). Weak error for nested multilevel Monte Carlo. Methodology and Computing in Applied Probability, 22(3):1325–1348.
- Lemaire, V. and Pagès, G. (2017). Multilevel Richardson-Romberg extrapolation.
- Pagès, G. (2007). Multi-step Richardson-Romberg extrapolation: remarks on variance control and complexity.
- Reisinger, C. and Wang, Z. (2018). Analysis of Multi-Index Monte Carlo estimators for a Zakai SPDE. Journal of Computational Mathematics, 36(2):202–236.
- Ruijter, M. and Oosterlee, C. (2012). Two-dimensional Fourier Cosine series expansion method for pricing financial options. SIAM Journal on Scientific Computing, 34(5):B642–B671.